### The homogeneity scale and the growth rate of cosmic structures

Felipe Avila,<sup>1</sup>\* Armando Bernui,<sup>1</sup> Rafael C. Nunes,<sup>2</sup> Edilson de Carvalho<sup>3</sup> and Camila P. Novaes<sup>1</sup>

<sup>1</sup>Observatório Nacional, Rua General José Cristino 77, São Cristóvão, 20921-400 Rio de Janeiro, RJ, Brazil
 <sup>2</sup>Instituto Nacional de Pesquisas Espaciais, Av. dos Astronautas 1758, Jardim da Granja, São José dos Campos, SP, Brazil
 <sup>3</sup>Centro de Estudos Superiores de Tabatinga, Universidade do Estado do Amazonas, 69640-000, Tabatinga, AM, Brazil

Accepted 2021 October 25. Received 2021 October 5; in original form 2021 June 14

### ABSTRACT

We propose a novel approach to obtain the growth rate of cosmic structures, f(z), from the evolution of the cosmic homogeneity scale,  $R_{\rm H}(z)$ . Our methodology needs two ingredients in a specific functional form:  $R_{\rm H}(z)$  data and the matter two-point correlation function today, i.e.  $\xi(r, z = 0)$ . We use a Gaussian Process approach to reconstruct the function  $R_{\rm H}$ . In the absence of suitable observational information of the matter correlation function in the local Universe,  $z \simeq 0$ , we assume a fiducial cosmology to obtain  $\xi(r, z = 0)$ . For this reason, our final result turns out to be a consistency test of the cosmological model assumed. Our results show a good agreement between: (i) the growth rate  $f^{R_{\rm H}}(z)$  obtained through our approach, (ii) the  $f^{\Lambda \text{CDM}}(z)$  expected in the fiducial model, and (iii) the best-fitting f(z) from data compiled in the literature. Moreover, using this data compilation, we perform a Gaussian Process to reconstruct the growth rate function  $f^{\text{data}}(z)$  and compare it with the function  $f^{R_{\rm H}}(z)$  finding a concordance of  $< 2\sigma$ , a good result considering the few data available for both reconstruction processes. With more accurate  $R_{\rm H}(z)$  data, from forthcoming surveys, the homogeneity scale function might be better determined and would have the potential to discriminate between  $\Lambda$ CDM and alternative scenarios as a new cosmological observable.

Key words: Cosmology: Observations - Cosmology: Large-Scale Structure of the Universe.

### **1 INTRODUCTION**

There is an increasing interest in measurements of the growth rate of cosmic structures, f(z), because this function behaves differently for cosmological models based on different theories of gravity (see e.g. Huterer et al. 2015; Kazantzidis & Perivolaropoulos 2018; Basilakos & Anagnostopoulos 2020; Linder 2020; Velasquez-Toribio & Fabris 2020); notoriously, the concordance cosmological model Lambda cold dark matter ( $\Lambda$ CDM) is based on the theory of general relativity. In this scenario, precise measurements of f(z) from diverse cosmological tracers measured at several redshifts would determine if the ACDM model correctly describes the evolution of the function f(z) (Pezzotta et al. 2017; Aubert et al. 2020; Bautista et al. 2021; Avila et al. 2021), and to investigate classes of models based on modified gravity theory (Alam et al. 2020; Ntelis et al. 2020). But the interest in f(z) is more fundamental. In fact, since the early works of Peebles (1965), Silk (1968), and Sunyaev & Zeldovich (1970), the theory of cosmological perturbations searches to describe the clustering evolution of the primordial density fluctuations, from the earliest times to the currently observed universe, where the growth rate of structures f(z) represents a measurement of such clustering evolution.

The measurements of f(z) can be done with good precision using the Redshift Space Distortions (RSD) approach, that is, studying the peculiar velocities caused by local gravitational potentials that introduce distortions in the two-point correlation function (2PCF; Kaiser 1987). Calculating the 2PCF from a galaxy survey, more precisely, the anisotropic correlation function,  $\xi(s, \mu)$  (Hamilton 1992; Hamilton & Culhane 1995), one can constrain the product  $f\sigma_8$ , where  $\sigma_8$  is the variance of the matter fluctuations at the scale of 8 Mpc/*h* (Juszkiewicz et al. 2009; Song & Percival 2009). For  $f\sigma_8$  data compilations, see e.g. Zhang & Li (2018), Sagredo, Nesseris & Sapone (2018), and Alam et al. (2021).

The growth rate of cosmic structures, f, is defined as (Strauss & Willick 1995)

$$f(a) \equiv \frac{\mathrm{d}\ln D(a)}{\mathrm{d}\ln a} \,,$$

where D = D(a) is the linear growth function, and *a* is the scale factor in the Robertson–Walker metric, based on general relativity theory. Apply the above equation to a catalogue of cosmic objects to measure *f* does not work, because what one can measure directly from the data survey is the density contrast  $\delta(r, a)$  and not the growth function D(a). In this work, we propose a solution for this problem: search for a cosmic observable function that depends only on cosmic time (equivalently, on the redshift *z* or the scale factor *a*) instead of D(a) in the above equation, being able to quantify the clustering evolution to provide a measurement of the growth rate of cosmic structures.

The Cosmological Principle is a fundamental piece in the concordance model of cosmology (Peebles 1980). It claims that, at sufficiently large scales, the universe is statistically homogeneous and isotropic (regarding the statistical isotropy of the universe, see

<sup>\*</sup> E-mail: felipeavila@on.br

e.g. Bernui et al. 2007; Bernui, Oliveira & Pereira 2014; Pereira & Pitrou 2015; Bengaly et al. 2017; Tarnopolski 2017; Dainotti, Del Vecchio & Tarnopolski 2018; Marques et al. 2018; Řípa & Shafieloo 2019). Several teams analysed galaxy surveys to calculate the scale where the transition from an inhomogeneous to a homogeneous distribution occurs, termed the homogeneity scale  $R_{\rm H}$  (Scrimgeour et al. 2012; Laurent et al. 2016; Ntelis et al. 2017). For recent analyses, see e.g. Ntelis et al. (2019), Heinesen (2020), Pandey (2021a), Pandey & Sarkar (2021b), Gonçalves et al. (2021), De Marzo, Labini & Pietronero (2021), and Camacho & Gaztañaga (2021). In addition, analyses of the angular scale homogeneity have also been done to find the angular scale of homogenenity,  $\theta_{\rm H}$ , (Alonso et al. 2015; Gonçalves et al. 2018a; Avila et al. 2018, 2019), considered model independent measurements because one does not assume a cosmological model, as in the analyses of  $R_{\rm H}$ , when one uses a fiducial cosmology to calculate 3D distances. At present, diverse deep astronomical surveys map large volumes of the universe, permitting to probe the evolution of  $R_{\rm H}$ , although it is not as accurate as desirable. The next generation of surveys foresees a large number of  $R_{\rm H}$  measurements with an improved accuracy (Amendola et al. 2018; Ivezić et al. 2019).

In this work, we will show that it is possible to use information from  $R_{\rm H}$ , more precisely from the homogeneity scale evolution  $dR_{\rm H}/dz$ , to obtain the cosmic evolution of the growth rate of structures f(z). From the theoretical point of view, the homogeneity scale can be related to the 2PCF,  $\xi(r)$  (Peebles 1980; Ntelis et al. 2017). From the linear perturbation theory, the redshift evolution of  $\xi(r)$ is proportional to  $D(z)^2$  then,  $F[R_{\rm H}(z)]D(z)^2 \propto$  cte, where *F* is a functional of the homogeneity scale function  $R_{\rm H}(z)$ . As we shall see, this proportionality leads to the growth rate f(z) through the redshift derivative of  $\xi(R_{\rm H}(z))$ , the volume-averaged 2PCF. The approach to know the functional *F* needs to assume parameters that we determine assuming a  $\Lambda$ CDM fiducial cosmology. In this sense, our analyses should be considered as a test of consistency for the  $\Lambda$ CDM model.

The relationship between f and  $R_{\rm H}$  indicates that with precise homogeneity scale data,  $R_{\rm H}(z)$ , measured at several redshifts, one can determine with good accuracy the growth rate of cosmic structures f = f(z), which in turn can be used to discriminate between the concordance  $\Lambda$ CDM and competing models based on modified gravity theories. Additionally, these data could be used in statistical analyses to find cosmological parameters. In other words, the homogeneity scale data,  $R_{\rm H}(z)$ , would indeed play the role of a novel cosmological observable, as first discussed by Ntelis et al. (2019).

This work is organized as follows. In Section 2, we review the main equations of the linear theory of matter perturbations. In Section 3, we explain the methodology to obtain the transition scale to homogeneity and, for the first time, the relation between  $R_{\rm H}(z)$  and f(z). In Section 4, we explain the reconstruction procedure to obtain a smooth curve of  $R_{\rm H}(z)$ , and  $dR_{\rm H}(z)/dz$ , both used then to obtain f(z) according to our procedure. In Sections 5, we show our results and discuss them, while in Section 6, we present our conclusions.

### **2 GROWTH RATE OF COSMIC STRUCTURES**

To describe the structure formation in an isotropic and homogeneous universe we used a perturbation approach: small deviation in the early universe has a slow evolution that can be described by a linear perturbation theory (Mukhanov, Feldman & Brandenberger 1992). One defines the density contrast as

$$\delta(\mathbf{r},t) \equiv \frac{\rho(\mathbf{r},t) - \bar{\rho}(t)}{\bar{\rho}(t)},\tag{1}$$

where  $\rho(\mathbf{r}, t)$  is the matter density at the comoving vector position  $\mathbf{r}$  at cosmic time t and  $\bar{\rho}(t)$  is the average matter density measured in the hyper-surface of constant t. In the linear and Newtonian regime, the gravitational potentials are small and the perturbation scale is smaller than the Hubble radius,  $\lambda \ll c/H_0$ , where c is the speed of light and  $H_0$  is the Hubble constant. Over this condition, the structure formation is described with the fluid equations

$$\dot{\delta} = -\frac{1}{a} \nabla \cdot \boldsymbol{v} \,, \tag{2}$$

$$\dot{\boldsymbol{v}} + H\boldsymbol{v} = -\frac{1}{a\bar{\rho}}\nabla\delta p - \frac{1}{a}\nabla\delta\Phi, \qquad (3)$$

$$\nabla^2 \delta \Phi = 4\pi G a^2 \bar{\rho} \,\delta\,,\tag{4}$$

which are the continuity, Euler, and Poisson equations, respectively, perturbed at first order in comoving space. The dot corresponds to a partial derivative in cosmic time. The physical quantities v,  $\delta p$ , and  $\delta \Phi$  are the peculiar velocity, pressure, and the perturbed gravitational potential, respectively.

Combining equations (2), (3), and (4), and assuming adiabatic perturbations condition, we obtain the well known equation that describes the linear density contrast evolution of the matter density

$$\ddot{\delta}_m + 2\frac{\dot{a}}{a}\delta_m - 4\pi G\bar{\rho}_m\delta_m = 0.$$
<sup>(5)</sup>

In the linear approximation, the density contrast is a function of time only, that is,  $\delta_m \sim D(t)$ . From this, one can define the growth rate of cosmic structures as

$$f(a) \equiv \frac{a}{D} \frac{\mathrm{d}D}{\mathrm{d}a} = \frac{\mathrm{d}\ln D}{\mathrm{d}\ln a} \,. \tag{6}$$

In the ACDM model we have the following approximation (Lahav et al. 1991)

$$f(z) \simeq \Omega_m^{0.6}(z) + \frac{\Omega_\Lambda}{70} \left( 1 + \frac{1}{2} \Omega_m(z) \right), \tag{7}$$

where  $\Omega_m$  and  $\Omega_{\Lambda}$  are the matter and dark energy cosmological parameters, respectively. An alternative approximation is given by (Linder 2005; Linder & Cahn 2007)

$$f(z) = \Omega_m^{\gamma}(z), \qquad (8)$$

where  $\gamma$  is the growth index. In the  $\Lambda$ CDM model,  $\gamma = 6/11 \simeq 0.55$ . This parameter assumes distinct values beyond  $\Lambda$ CDM cosmology (Basilakos 2012).

### **3 TRANSITION SCALE TO HOMOGENEITY**

The most used methodology to study the homogeneity of galaxy or quasars distributions is to count the number of cosmic objects,  $N_{gal}$ , inside a sphere of radius *r*, and divide for  $N_{rand}$ , the equivalent count but for a random distribution, that has the same features as the original one. Then, we can define the *scaled counts-in-spheres*,  $\mathcal{N}(< r)$  (Scrimgeour et al. 2012)

$$\mathcal{N}(< r) \equiv \frac{N_{\text{gal}}(< r)}{N_{\text{rand}}(< r)},\tag{9}$$

where for a homogeneous distribution, at large scales, it goes to 1. It can be shown that  $\mathcal{N}(< r)$  is related to the two-point correlation

function  $\xi(r)^1$ 

$$\mathcal{N}(< r) = 1 + \frac{3}{r^3} \int_0^r \xi(s) s^2 \mathrm{d}s \,. \tag{10}$$

From the function  $\mathcal{N}(< r)$ , one can define the correlation dimension function  $\mathcal{D}_2(r)$  (for details, see the appendix A in Ntelis et al. 2017),

$$\mathcal{D}_2(r) \equiv \frac{\mathrm{d}\,\ln\mathcal{N}(< r)}{rmd\,\ln r} + 3\,. \tag{11}$$

Despite the fact that most studies present both estimators,  $\mathcal{N}(< r)$  and  $\mathcal{D}_2(r)$ , the result from the correlation dimension is considered more robust, because it is less correlated for most scales (Scrimgeour et al. 2012; Ntelis et al. 2017).

To finish this section, we discuss the arbitrary criterion used to determine the scale where the transition to homogeneity occurs,  $R_{\rm H}$ . Consider the following equation

$$\mathcal{D}_2(R_{\rm H}) = 3(1 - \epsilon). \tag{12}$$

In an ideal situation, one expects  $\epsilon = 0$ , that is, when the homogeneity scale is attained the value of  $\epsilon$  should be  $\epsilon = 0$ , such that the transition to homogeneity occurs on the scale at which  $D_2$  calculated from data achieve the value 3. However, due to systematic effects present in the galaxy surveys, Scrimgeour et al. (2012) suggested to fix the value of  $\epsilon$  at, for example, 0.01, which gives us  $D_2(R_{\rm H}) = 2.97$ , that is, 1 per cent below 3. We assume this value because it is commonly adopted in the literature, and allow us to study the scale of homogeneity for different tracers in a large range of redshift. Anyhow, as we will show next, our methodology is independent of  $\epsilon$ , due to the redshift derivative.

## 3.1 The growth rate of cosmic structures from the homogeneity scale

As mentioned above, the scaled counts-in-spheres,  $\mathcal{N}(< r)$ , is related to the two-point correlation function,  $\xi(r; z)$ , at redshift z

$$\xi(r = |\mathbf{x} - \mathbf{y}|; z) = \langle \delta(\mathbf{x}; z) \delta(\mathbf{y}; z) \rangle, \qquad (13)$$

that is, is the spatial average of the product of the density contrasts evaluated at the arbitrary positions of a pair of galaxies, x, y, at redshift z. The redshift evolution of  $\xi$  can be obtained assuming for the equation (5) the solution  $\delta(r; z) = \delta(r; z = 0)D(z)$  (Schneider 2006). This leads to

$$\xi(r = |\mathbf{x} - \mathbf{y}|; z) = \langle \delta(\mathbf{x}; z) \delta(\mathbf{y}; z) \rangle$$
  
=  $D^2(z) \langle \delta(\mathbf{x}; z = 0) \delta(\mathbf{y}; z = 0) \rangle$   
=  $D^2(z) \xi(r; z = 0)$ , (14)

where  $\xi(r; z = 0)$  is the two-point correlation function at z = 0. From equation (14), one can rewrite the scaled counts-in-spheres as

$$\mathcal{N}(\langle r, z) = 1 + D^2(z)\bar{\xi}(r),$$
 (15)

where

$$\bar{\xi}(r) \equiv \frac{3}{r^3} \int_0^r \xi(s, z=0) s^2 \mathrm{d}s$$
(16)

is the volume average of the correlation function. From the equation (11), one has

$$\mathcal{D}_2(r,z) = \frac{rD^2(z)}{1+D^2(z)\bar{\xi}(r)} \frac{\mathrm{d}\xi(r)}{\mathrm{d}r} + 3.$$
(17)

<sup>1</sup>For applications of the two-point correlation function in clustering analyses, see e.g. de Carvalho et al. (2018, 2021) and Carvalho et al. (2020).

It is useful to define the following quantity:

$$\zeta(r) \equiv \frac{\mathrm{d}\xi(r)}{\mathrm{d}r} \,. \tag{18}$$

For the scale where the transition to homogeneity occurs,  $r = R_{\rm H}$ , equation (17) becomes

$$R_{\rm H}(z)D^2(z)\zeta[R_{\rm H}(z)] = -3\epsilon \left(1 + D^2(z)\bar{\xi}[R_{\rm H}(z)]\right) \simeq -3\epsilon , \qquad (19)$$

where we consider only the first-order term. Now, taking the redshift derivative of equation (19) we have

$$\frac{\mathrm{d}}{\mathrm{d}z}\left(R_{\mathrm{H}}D^{2}\zeta\right)=0\,.\tag{20}$$

This differential equation relates in a simple way  $R_{\rm H}(z)$  and D(z). Finally, separating each term of equation (20), we have

$$-\frac{2}{D}\frac{\mathrm{d}D}{\mathrm{d}z} = \frac{1}{R_{\mathrm{H}}}\frac{\mathrm{d}R_{\mathrm{H}}}{\mathrm{d}z} + \frac{1}{\zeta}\frac{\mathrm{d}\zeta}{\mathrm{d}z}.$$
(21)

Using the equation (6) in equation (21), we have

$$f(z) = \frac{1+z}{2} \left( \frac{1}{R_{\rm H}} \frac{\mathrm{d}R_{\rm H}}{\mathrm{d}z} + \frac{1}{\zeta} \frac{\mathrm{d}\zeta}{\mathrm{d}z} \right),\tag{22}$$

which, explicitly, is independent of  $\epsilon$ . To obtain f(z), in addition to  $R_{\rm H}$  data, we must obtain  $\zeta [R_{\rm H}(z)]$  from a correlation function in z = 0. Or, in a model dependent way, use an approximation, as we will describe in the next section.

### **4 TESTING THE MODEL**

In this section, we describe our methodology, aimed to solve equation (22), and apply it to a set of  $R_{\rm H}$  data to obtain the growth function f(z). First, we present the data, and secondly, we detail the approximation used to define  $\zeta(R_{\rm H})$ . By last, we describe the Gaussian Process methodology used to reconstruct  $R_{\rm H}$  and  $dR_{\rm H}/dz$ .

### 4.1 Data

Here, we use two sets of  $R_{\rm H}$  measurements. The first one is provided by Ntelis et al. (2017), through the study of the CMASS galaxy sample of the BOSS survey, they calculated the transition to homogeneity for 5 uncorrelated redshift bins in the interval 0.43–0.70. The authors analysed separately the North (NGC) and South Galactic Caps (SGC), at the same redshift bins, obtaining five independent measurements for each of them (i.e. a total of 10  $R_{\rm H}$  data). The second  $R_{\rm H}$  data set comes from Gonçalves et al. (2018b), who analysed the quasars sample from the fourteenth data release of the Sloan Digital Sky Survey (SDSS-IV DR14) in the redshift interval 0.80–2.24. They measured  $R_{\rm H}$  in each one of four uncorrelated redshift bins (i.e. 4  $R_{\rm H}$  data) studied employing two estimators to calculate  $\mathcal{N}(< r)$ : Landy-Szalay (LS) and Peebles–Hauser (PH) estimators, obtaining similar results in both cases. In Table 1, we list these  $R_{\rm H}(z)$  data with their respective redshifts.

Following Ntelis et al. (2019), we combine the  $R_{\rm H}$  data from NGC (*j*) and SGC (*k*) using a weighted average, defined as

$$R_{\rm H}^{\rm w}(z_i) \equiv \left(\frac{1}{\sigma_j^2(z_i)} + \frac{1}{\sigma_k^2(z_i)}\right)^{-1} \times \left(\frac{R_{\rm H}^j(z_i)}{\sigma_j^2(z_i)} + \frac{R_{\rm H}^k(z_i)}{\sigma_k^2(z_i)}\right), \quad (23)$$

where  $\sigma_j$ ,  $\sigma_k$ , and  $R_{\rm H}^j$ ,  $R_{\rm H}^k$  are, respectively, the errors and data for two independent measurements, *j* and *k*, in the same redshift,  $z_i$ . At first order, we can neglect the covariance between redshift bins. For the Gonçalves et al. (2018b) data, we choose the data from the PH estimator to optimize our analyses.



**Figure 1.** Left-hand panel: Relative error between  $\zeta[R_H]$  as fitted using the DPL approximation (equation 24) and the expectation from the  $\Lambda$ CDM model. Right-hand panel: Relative error of the cosmic growth rate f(z) calculated using equation (26) with respect to f(z) obtained using the equation (6) assuming the same  $\Lambda$ CDM fiducial cosmology; one also observes the noisy pattern caused by the numerical derivative.

Notice that the  $R_{\rm H}$  we use here are already corrected by the corresponding bias factor. The bias for each redshift bin have been determined by the respective authors, which provided the bias-corrected measurements. This is important because each homogeneity scale measurement is calculated for a specific tracer and a proper combination of these data requires their conversion to the corresponding transition scale for the underlying matter distribution.

### 4.2 Double power-law approximation for $\zeta(R_{\rm H})$

In order to use the selected  $R_{\rm H}$  data sample to calculate  $f^{R_{\rm H}}(z) = f(z)$ , using equation (22), we need to define  $\zeta(R_{\rm H})$ . For this we approximate  $\zeta(R_{\rm H})$  by a Double Power Law (DPL) function, similar to that one used in the study of the AGN luminosity function (Kulkarni, Worseck & Hennawi 2019)

$$\zeta(R_{\rm H}) = -\frac{C R_{\rm H}^{-1}}{(R_{\rm H}/R_{\star})^{\alpha} + (R_{\rm H}/R_{\star})^{\beta}},$$
(24)

where *C*,  $R_{\star}$ ,  $\alpha$ , and  $\beta$  are the parameters to be adjusted. Taking its redshift derivative

$$\frac{1}{\zeta} \frac{d\zeta}{dz} = -\frac{(1+\alpha)(R_{\rm H}/R_{\star})^{\alpha} + (1+\beta)(R_{\rm H}/R_{\star})^{\beta}}{(R_{\rm H}/R_{\star})^{\alpha} + (R_{\rm H}/R_{\star})^{\beta}} \frac{1}{R_{\rm H}} \frac{dR_{\rm H}}{dz}, \quad (25)$$

the growth rate can be written as

$$f^{R_{\rm H}}(z) = \frac{1+z}{2} \left[ 1 - \frac{\left(1+\alpha\right) \left(\frac{R_{\rm H}}{R_{\star}}\right)^{\alpha} + \left(1+\beta\right) \left(\frac{R_{\rm H}}{R_{\star}}\right)^{\beta}}{\left(\frac{R_{\rm H}}{R_{\star}}\right)^{\alpha} + \left(\frac{R_{\rm H}}{R_{\star}}\right)^{\beta}} \right] \frac{1}{R_{\rm H}} \frac{\mathrm{d}R_{\rm H}}{\mathrm{d}z} ,$$
(26)

where  $R_{\rm H} = R_{\rm H}(z)$ .

We fit the four free parameters of the DPL approximation to the theoretical expectation for  $\zeta(R_{\rm H})$ , for  $R_{\rm H}$  corresponding the redshift range 0 < z < 2, considering the  $\Lambda$ CDM model baseline obtained from Planck Collaboration (2020), that is, h = 0.6727,  $\Omega_c h^2 = 0.1202$ ,  $\Omega_b h^2 = 0.02236$ ,  $\Sigma m_v = 0.0600$ ,  $n_s = 0.9649$ ,  $\sigma_8 = 0.8120$ , and  $\ln(10^{10}A_s) = 3.045$ . For this we employ the public code cosmopit<sup>2</sup> (Ntelis et al. 2017, 2018) to produce  $\zeta(R_{\rm H})^{\Lambda \text{CDM}}$ , which uses the public code CLASS<sup>3</sup> (Lesgourgues 2011; Blas, Lesgourgues & Tram 2011) as a background. We obtain for these parameters  $[R_\star, \alpha, \beta, C] = [46.16, 2.76, 1.12, 0.19]$ , whose

<sup>2</sup>https://github.com/lontelis/cosmopit

<sup>3</sup>https://github.com/lesgourg/class\_public

**Table 1.** The  $R_{\rm H}(z)$  data used in the analyses.

z	$R_{\rm H}$ (M	pc $h^{-1}$ )	Reference	
	NGC	SGC		
0.457	$64.20 \pm 1.30$	$66.70 \pm 1.60$	Ntelis et al. (2017)	
0.511	$65.40 \pm 0.90$	$63.90 \pm 1.50$		
0.565	$62.60\pm0.80$	$65.20 \pm 1.60$		
0.619	$60.40 \pm 0.80$	$60.10 \pm 1.10$		
0.673	$59.00\pm0.80$	$60.10 \pm 1.80$		
	PH	LS		
0.985	$48.78\pm3.82$	$52.93 \pm 7.55$	Gonçalves et al. (2018b)	
1.350	$40.56 \pm 3.39$	$40.43 \pm 5.64$		
1.690	$36.19 \pm 3.45$	$36.66 \pm 4.80$		
2.075	$27.91 \pm 3.91$	$29.94\pm3.35$		

error for each of them is less than 1 per cent. The plot on the lefthand panel of Fig. 1 shows the relative error between the input ACDM expectation and the fitted DPL approximation, where we observe a good agreement on all scales. The maximum discrepancy of 0.3 percent appears at the largest  $R_H$  scales considered here. This occurs due to the effect introduced in  $\zeta(R_H)$  by the presence of the BAO feature at  $\sim 100h^{-1}$ Mpc in the correlation function, then the DPL approximation fails to model the large scales. See Appendix A for more details. Notice that, although the methodology does not depend on  $\epsilon$  explicitly, as shown in equation (22), we follow Scrimgeour et al. (2012) and fix this value to  $\epsilon = 0.01$  to obtain the  $\zeta [R_{\rm H}(z)]^{\Lambda \rm CDM}$  function and then calculate the best-fitting parameters for the DPL approximation. In Appendix B, we test the criterion for  $\epsilon$  and the dependence of our methodology on some cosmological parameters. Additionally, the right panel of Fig. 1 shows the relative error between the cosmic growth rate  $f^{R_{\rm H}}(z)$ calculated using the DPL approximation equation (26) and that obtained from equation (6) assuming the ACDM fiducial model. As observed, for almost the whole redshift interval, we observe a maximum deviation of  $\sim 0.5$  per cent, again at low redshifts, where we also notice the noisy pattern caused by the numerical derivative. Note that, since we have a small set of  $R_{\rm H}$  data, we use a reconstructed function from them to be able to appropriately calculate the derivative  $dR_{\rm H}/dz$  in equation (26), a procedure detailed in the following section.

#### 4.3 Gaussian process regression

To extract maximum cosmological information from the  $R_{\rm H}$  data listed in Table 1 we perform a Gaussian Process (GP) Regression



**Figure 2.** Left-hand panel: Reconstruction of the homogeneity scale function  $R_{\rm H}(z)$  using Gaussian Process (dashed line) and the  $R_{\rm H}$  measurements (red squares) presented in Table 1 (see the text for details about the data set); the shadow represents the 95 per cent CL. Right-hand panel: Derivation of the growth rate of structures  $f^{R_{\rm H}}(z)$  (dashed line) using the reconstructed function  $R_{\rm H}(z)$ , shown in the left-hand panel plot, and the equation (26), where the shadow represents the 95 per cent CL. The solid line represents the expression expected in the  $\Lambda$ CDM model,  $f^{\Lambda \text{CDM}}(z)$ , obtained from equation (6) with  $\Omega_{\rm Planck}^{\rm Planck} = 0.315$ ; instead, the dot–dashed line shows the expected growth rate using equation (6) but with  $\Omega_m = 0.270^{+0.073}_{-0.073}$ , a value obtained from the best-fitting analyses of the f(z) data as shown in Fig. 4 (the red squares are the data listed in Table D1). Comparing the  $f^{R_{\rm H}}(z)$  (dashed line) and the best-fitting f(z) (dot–dashed line) functions we found an agreement of  $< 2\sigma$  (considering the corresponding uncertainties, not shown in the figure to avoid excess of information).

method, obtaining in this way a smooth curve for  $R_{\rm H}(z)$  and, by numerical derivation, for  $dR_H/dz$ ; this information is then used in equation (26) to obtain the  $f^{R_{\rm H}}(z)$  function. The GP consists of generic supervised learning method designed to solve regression and probabilistic classification problems, where we can interpolate the observations and compute empirical confidence intervals and a prediction in some region of interest (Rasmussen 2003; Pezzotta et al. 2017). The GP method design from machine learning techniques is the state of the art to obtain statistical information and model prediction from some previously known information or data. In the cosmological context, GP techniques has been used to reconstruct cosmological parameters, like the dark energy equation of state,  $\omega(z)$ , the expansion rate of the universe, the cosmic growth rate, and other cosmological functions (see e.g. Seikel, Clarkson & Smith 2012; Shafieloo, Kim & Linder 2012; Zhang & Li 2018; Marques et al. 2019; 2020; ; Nunes et al. 2020a; Nunes & Bernui 2020b; Renzi & Silvestri 2020; Bonilla, Kumar & Nunes 2021a; Bonilla et al. 2021b; Colgáin & Sheikh-Jabbari 2021; Escamilla-Rivera, Said & Mifsud 2021; Sun, Jiao & Zhang 2021 for a short list of references).

The main advantage in this procedure is that it is able to make a non-parametric inference using only a few physical considerations and minimal cosmological assumptions. Our aim is to reconstruct a function  $F(x_i)$  from a set of its measured values  $F(x_i) \pm \sigma_i$ , where  $x_i$  represent our data sample. It assumes that the value of the function at any point  $x_i$  follows a Gaussian distribution. The value of the function at  $x_i$  is correlated with the value at other point  $x'_i$ . Thus, a GP is defined as

$$F(x_i) = \mathcal{GP}(\mu(x_i), \operatorname{cov}[F(x_i), F(x_i)]),$$
(27)

where  $\mu(x_i)$  and  $\operatorname{cov}[F(x_i), F(x_i)]$  are the mean and the variance of the variable at  $x_i$ , respectively. For the reconstruction of the function  $F(x_i)$ , the covariance between the values of this function at different positions  $x_i$  can be modeled as

$$cov[F(x), F(x')] = k(x, x'),$$
 (28)

where k(x, x') is known as the kernel function. The kernel choice is often crucial for obtaining good results regarding the reconstruction of the function of interest.

The kernel most commonly used is the standard Gaussian Squared-Exponential approach, which is defined as

$$k(x, x') = \sigma_F^2 \exp\left(-\frac{|x - x'|^2}{2l^2}\right),$$
(29)

where  $\sigma_F^2$  is the signal variance, which controls the strength of the correlation of the function *F*, and *l* is the length scale that determines the capacity to model the main characteristics (global and local) in the evaluation region to be predicted (or coherence length of the correlation in *x*). These two parameters are often called hyperparameters.

It is well known that depending on the data set in analysis, the kernel choice is an important point. We verify that our data set is well modelled by the choice above, and that other kernels do not produce major changes in our main results (see Appendix B). In what follows, we discuss our results.

### **5 RESULTS AND DISCUSSIONS**

In this section, we present our main results. We performed a GP to reconstruct the homogeneity scale,  $R_H(z)$ . From this, using the DPL model, we can obtain  $f^{R_H}(z)$ . Also, we perform the GP to our f(z) data compilation. Finally, we study the parameter space  $H_0 - \Omega_m$  from the same data compilation.

### 5.1 Results of the reconstruction of $f^{R_{\rm H}}(z)$ and f(z)

In obtaining the results to be described in this section, we use the Scikit-learn code (Pedregosa et al. 2011), which is a PYTHON module integrating a wide range of state-of-the-art machine learning algorithms, to model the GP described in the previous section. The hyperparameters  $\sigma_F^2$  and *l* are optimized during the fitting by maximizing the log-marginal-likelihood.

The left-hand panel of Fig. 2 shows the best-fitting prediction of the reconstruction GP of the homogeneity scale function  $R_H(z)$ , at 95 per cent CL, from the data sample listed in Table 1 (represented by red squares in this plot). As verified in this plot, the  $R_H(z)$  data reveals the expected behaviour in the evolution of the matter clustering in



**Figure 3.** Comparison of the growth rate  $f^{R_{\rm H}}(z)$  obtained from equation (26) (dashed line) and the GP reconstructed function  $f^{\rm data}(z)$  (dot–dashed line) using the data compilation given in Table D1. Both reconstructed functions show a significant overlapping of the respective  $2\sigma$  regions (shaded areas).



**Figure 4.** The 68 per cent (dark shaded area) and 95 per cent (light shaded area) CLs regions, respectively, on the parametric space  $H_0 - \Omega_m$  from f(z) + Planck- $H_0$  prior and f(z) + BAO joint analyses. The parameter  $H_0$  is measured in the units of km s<sup>-1</sup> Mpc<sup>-1</sup>.

the Universe, going from a nearly homogeneous situation at high redshift to a non-linear clustered matter at low redshift where the homogeneity scale is attained only at large scales.

We use these  $R_H(z)$  data to obtain the evolution of the growth rate of cosmic structures,  $f^{R_H}(z)$ , according to equation (26) following the procedure described in Section 3.1. Our result can be observed on the right panel of Fig. 2, where  $f^{R_H}(z)$  is plotted as a dashed line and the current measurements of f(z), listed in the Table D1 in appendix D, as red squares. It is important to mention that  $f^{R_H}$  obtained through our procedure does not represent a direct f(z) measurement, but a non-parametric inference that can describe the evolution of the growth rate function from minimal cosmological assumptions. We also show for comparison the  $\Lambda$ CDM expected growth rate using equation (6) in two cases: using  $\Omega_m^{\text{Planck}} = 0.315$  (continuous line) from the Planck cosmological parameters, and using  $\Omega_m = 0.26$ (dot–dashed line) from the best-fitting data analyses shown in Fig. 4. On the other hand, it is interesting to compare the growth rate of cosmic structures  $f^{R_{\rm H}}(z)$  from the evolution of the cosmic homogeneity scale, with the  $f^{\rm data}(z)$  resulting from a GP reconstruction using the current f(z) data listed in Table D1. Notice that, the reconstruction procedure of  $f^{R_{\rm H}}(z)$  is performed in the redshift interval with  $R_{\rm H}(z)$  data, namely  $z \in [0.457, 2.075]$ , while the reconstruction procedure of  $f^{\rm data}(z)$  is done with f(z) data in the interval  $z \in [0.013, 1.4]$ . Then, for the comparative analysis we consider the common redshift interval:  $z \in [0.457, 1.4]$  shown in Fig. 3, where we observe that both functions agree well ( $<2\sigma$  level), overlapping significantly.

One should notice that a plausible systematic present in the  $R_{\rm H}$  data, listed in Table 1, is sourced by the necessity to assume a fiducial cosmology to calculate the 3D distances to the cosmic objects (galaxies or quasars), so that one can determine the 3D separation distance between each pair of them, information used to measure  $R_{\rm H}$ . As a matter of fact, the  $R_{\rm H}(z)$  measurements are model dependent and one should be cautious with this. For instance, the analyses done by Ntelis et al. (2017) assumed a fiducial cosmology different to that assumed by Gonçalves et al. (2018b), a fact that helps to explain why in the left panel of Fig. 2 one data set appear slightly over and the other slightly under the reconstructed function (dashed line).

### **5.2** Validation test of $H_0 - \Omega_m$ plane estimates from the current compilation of growth rate data

As a final discussion of this section, we will check what our compilation of f(z) data, shown in Table D1, can tell us about the  $\Lambda$  CDM baseline. Let us perform an analysis in three steps:

(i) To constrain  $\Omega_m$  we consider f(z) data (see Table D1) only.

(ii) A combination f(z) data plus a Gaussian prior on  $H_0$  using the Planck-CMB best fit. Note that we are within a  $\Lambda$ CDM baseline, so to use Planck-CMB information in  $\Lambda$ CDM itself context is just to improve the constraint on  $\Omega_m$ .

(iii) We consider the joint analysis f(z) + BAO. In this work, we consider the most recent BAO data compilation comprised of the  $D_V(z)/r_d$ ,  $D_M(z)/r_d$ , and  $D_H(z)/r_d$  measurements compiled in Table 3 in Alam et al. (2021).

We use the Metropolis-Hastings mode in CLASS + MontePython code (Lesgourgues 2011; Blas et al. 2011; Audren et al. 2013; Brinckmann & Lesgourgues 2019) to derive the constraints on cosmological parameters from the data sets described above, ensuring a Gelman–Rubin convergence criterion of  $R - 1 < 10^{-3}$ .

Fig. 4 shows the parameter space in the  $H_0 - \Omega_m$  plane at 68 per cent and 95 per cent CL from f(z) + Planck- $H_0$  prior and f(z) + BAO joint analyses, where f(z) data refers to the measurements presented in Table D1. The summary of the main results of our statistical analyses at 68 per cent CL are:  $\Omega_m = 0.27^{+0.079}_{-0.073}$  (f(z) data only),  $\Omega_m = 0.279^{+0.067}_{-0.067}$  (f(z) +  $H_0$ -Planck) and  $\Omega_m = 0.291^{+0.033}_{-0.030}$  and  $H_0 = 67.4^{+2.1}_{-2.0}$  km s<sup>-1</sup> Mpc<sup>-1</sup> from f(z) + BAO combination.

As well known, there is a growing tension for low z measurements of f(z) and it is weaker than the Planck- $\Lambda$ CDM predictions (see Di Valentino et al. 2021; Perivolaropoulos & Skara 2021 and reference therein for a review and the recent discussion presented in Nunes & Vagnozzi 2021). Our results here also confirm that growth rate data based on the measurements in Table D1 predict a suppression on the amplitude of the matter density perturbation at low z due the low  $\Omega_m$ estimation in comparison with that from the Planck- $\Lambda$ CDM,  $\Omega_m =$ 0.315 ± 0.007 (Planck Collaboration 2020). On the right-hand panel of Fig. 2, we also show the theoretical curve assuming our constraint on  $\Omega_m$ . Despite predicting a low  $\Omega_m$  best-fitting value in our analysis, the error bar estimates are in agreement with Planck CMB at  $< 1\sigma$ .

#### **6** CONCLUSIONS

Measurements of the growth rate of cosmic structures, f, have the potential to differentiate between the theory of general relativity, that supports the concordance model  $\Lambda$ CDM, from alternative scenarios based on modified gravity models. Besides the efforts, the current uncertainties in such measurements do not allow to discern between competing models of modified gravity.

This motivated us to search for a cosmological observable that depends only on the cosmic time *t*, or equivalently on the redshift *z*. We propose to use the transition scale to homogeneity,  $R_{\rm H}(z)$ , to know the evolution of the growth rate f = f(z). As shown in the Section 3.1, the relation between  $R_{\rm H}$  and *f* is not direct and one needs two ingredients: (i) a set of  $\{R_{\rm H}(z_i)\}$  data – in the redshift interval of interest – to reconstruct the continuous function  $R_{\rm H}(z)$  and to perform its redshift derivative; and (ii) the matter two-point correlation function at z = 0,  $\xi(r, z = 0)$ , that analyses distance scales of the order of the homogeneity scale. However, there is no observational data to construct  $\xi(r, z = 0)$ , and one has to assume a fiducial cosmology to obtain it. For this reason, our analyses and results are actually consistency tests of the cosmological model assumed.

Using GP, our reconstruction of the homogeneity scale function  $R_{\rm H}(z)$  done in Section 5 shows the expected behaviour, although the current data set is small and with large errors (see Table 1). With the functions  $R_{\rm H}(z)$  and  $\xi(r, z = 0)$ , and following our procedure, we use them in equation (26) to obtain the growth rate of cosmic structures  $f^{R_{\rm H}}(z)$ . Our results, displayed in the right-hand panel of fig. 2, show a good agreement between: (i) the growth rate  $f^{R_{\rm H}}(z)$  obtained through our approach, (ii) the  $f^{\Lambda \rm CDM}(z)$  expected in the fiducial model, and (iii) the best-fit f(z) from the set of  $\{f(z_i)\}$  measurements available in the literature. Moreover, using this compilation of  $\{f(z_i)\}$  data (see Appendix D), we perform a GP to reconstruct the growth rate function  $f^{\rm data}(z)$  and compare it with the function  $f^{R_{\rm H}}(z)$  obtained from our approach, finding a concordance of  $< 2\sigma$  as observed in Fig. 3 (notice the significant overlapping of their  $2\sigma$  regions). This is a good result considering the few data available for both reconstruction processes.

It is worth to note that our approach to find the growth rate of cosmic structures, f(z), from the evolution of the cosmic homogeneity scale,  $R_{\rm H}(z)$ , relies on the definition of the homogeneity scale which is not unique (see e.g. Pandey 2021a; Pandey & Sarkar 2021b); in our approach, the homogeneity scale is provided by the estimator  $D_2$  through the analyses of the universe fractal structure (Scrimgeour et al. 2012).

The relationship found between *f* and  $R_{\rm H}$  indicates that with precise homogeneity scale data,  $R_{\rm H}(z)$ , measured at several redshifts from forthcoming surveys (see e.g. Amendola et al. 2018; Ivezić et al. 2019), one can determine with good accuracy the growth rate of cosmic structures f = f(z), which in turn can be used to discriminate between the concordance  $\Lambda$ CDM and competing models based on modified gravity theories. Moreover, these data could be used in statistical analyses to find cosmological parameters. In summary, the homogeneity scale data,  $R_{\rm H}(z)$ , would indeed play the role of a novel cosmological observable, as first discussed by Ntelis et al. (2019).

### ACKNOWLEDGEMENTS

FA, AB, EdC, and CPN thank CAPES, CNPq, PROPG-CAPES/FAPEAM program, and FAPESP (process no. 2019/06040-0) for the grants under which this work was carried out. RCN

acknowledges financial support from the Fundação de Amparo à Pesquisa do Estado de São Paulo (FAPESP; São Paulo Research Foundation) under the project no. 2018/18036-5.

### DATA AVAILABILITY

The data underlying this article will be shared on request to the corresponding author.

### REFERENCES

- Alam S. et al., 2020, preprint (arXiv:2011.05771)
- Alam S. et al., 2021, Phys. Rev. D, 103, 083533
- Alonso D. et al., 2015, MNRAS, 449, 670
- Amendola L. et al., 2018, Living Rev. Rel., 21, 2
- Aubert M. et al., 2020, preprint (arXiv:2007.09013)
- Audren B. et al., 2013, J. Cosmol. Astropart. Phys., 02, 001
- Avila F. et al., 2018, J. Cosmol. Astropart. Phys., 12, 041
- Avila F. et al., 2019, MNRAS, 488, 1481
- Avila F. et al., 2021, MNRAS, 505, 3404
- Basilakos S., 2012, Int. J. Mod. Phys. D, 21, 1250064
- Basilakos S., Anagnostopoulos F. K., 2020, Eur. Phys. J. C, 80, 1
- Bautista J. E. et al., 2021, MNRAS, 500, 736
- Bengaly C. A. P., Bernui A., Ferreira I. S., Alcaniz J. S., 2017, MNRAS, 466, 2799
- Bernui A., Mota B., Rebouças M. J., Tavakol R., 2007, Int. J. Mod. Phys. D, 16, 411
- Bernui A., Oliveira A. F., Pereira T. S., 2014, J. Cosmol. Astropart. Phys., 10, 041
- Blake C. et al., 2011, MNRAS, 415, 2876
- Blake C. et al., 2013, MNRAS, 436, 3089
- Blas D., Lesgourgues J., Tram T., 2011, J. Cosmol. Astropart. Phys., 07, 034
- Bonilla A. et al., 2021, preprint (arXiv:2102.06149)
- Bonilla A., Kumar S., Nunes R. C., 2021, Eur. Phys. J. C, 81, 1
- Brinckmann T., Lesgourgues J., 2019, Phys. Dark Universe, 24, 100260
- Camacho B., Gaztañaga E., 2021, preprint (arXiv:2106.14303)
- Carvalho G. C. et al., 2020, Astropart. Phys., 119, 102432

Colgáin E. Ó., Sheikh-Jabbari M. M., 2021, preprint (arXiv:2101.08565)

- Da Ângela J. et al., 2008, MNRAS, 383, 565
- Dainotti M. G., Del Vecchio R., Tarnopolski M., 2018, Adv. Astron., 2018, 4969503
- de Carvalho E. et al., 2018, J. Cosmol. Astropart. Phys., 04, 064
- de Carvalho E. et al., 2021, A&A, 649, A20
- De Marzo G., Labini F. S., Pietronero L., 2021, A&A, 651, A114
- Di Valentino E. et al., 2021, Astropart. Phys., 131, 102604
- Eisenstein D. J., Hu W., 1998, ApJ, 496, 605
- Escamilla-Rivera C., Said J. L., Mifsud J., 2021, J. Cosmol. Astropart. Phys., 10, 016
- Gonçalves R. S., 2021, J. Cosmol. Astropart. Phys., 03, 029
- Gonçalves R. S. et al., 2018a, MNRAS, 475, L20
- Gonçalves R. S. et al., 2018b, MNRAS, 481, 5270
- Guzzo L., 2008, Nature, 451, 541
- Hamilton A. J. S., 1992, ApJ, 385, L5
- Hamilton A. J. S., Culhane M., 1995, MNRAS, 278, 73
- Hawkins E. et al., 2003, MNRAS, 346, 78
- Heinesen A., 2020, J. Cosmol. Astropart. Phys., 10, 052
- Huterer D. et al., 2015, Astropart. Phys., 63, 23
- Ivezić Ž. et al., 2019, ApJ, 873, 111
- Juszkiewicz R. et al., 2009, J. Cosmol. Astropart. Phys., 2, 021 Kaiser N., 1987, MNRAS, 227, 1
- Kazantzidis L., Perivolaropoulos L., 2018, Phys. Rev. D, 97, 103503
- Kulkarni G., Worseck G., Hennawi J. F., 2019, MNRAS, 488, 1035
- Lahav O. et al., 1991, MNRAS, 251, 128
- Laurent P., 2016, J. Cosmol. Astropart. Phys., 11, 060
- Lesgourgues J., 2011, preprint (arXiv:1104.2932)
- Linder E. V., 2005, Phys. Rev. D, 72, 043529
- Linder E. V., 2020, J. Cosmol. Astropart. Phys., 10, 042

- Linder E. V., Cahn R. N., 2007, Astropart. Phys., 28, 481
- Marques G. A., Novaes C. P., Bernui A., Ferreira I. S., 2018, MNRAS, 473, 165
- Marques G. A. et al., 2019, J. Cosmol. Astropart. Phys., 06, 019
- Marques G. A., Liu J., Huffenberger K. M., Colin Hill J., 2020, ApJ, 904, 182
- Mukhanov V. F., Feldman H. A., Brandenberger R. H., 1992, Phys. Rep., 215, 203
- Ntelis P. et al., 2017, J. Cosmol. Astropart. Phys., 06, 019
- Ntelis P. et al., 2018, J. Cosmol. Astropart. Phys., 12, 014
- Ntelis P. et al., 2019, preprint (arXiv:1904.06135)
- Ntelis P. et al., 2020, preprint (arXiv:2010.06707)
- Nunes R. C., Bernui A., 2020b, Eur. Phys. J. C, 80, 1025
- Nunes R. C., Vagnozzi S., 2021, MNRAS, 505, 5427
- Nunes R. C. et al., 2016, J. Cosmol. Astropart. Phys., 08, 051
- Nunes R. C., Yadav S. K., Jesus J. F., Bernui A., 2020a, MNRAS, 497, 2133
- Pandey B., 2021, J. Cosmol. Astropart. Phys., 02, 023
- Pandey B., Sarkar S., 2021, J. Cosmol. Astropart. Phys., 07, 019
- Pedregosa F. et al., 2011, J. Mach. Learn. Res., 12, 2825
- Peebles P. J. E., 1965, ApJ, 142, 1317
- Peebles P. J. E., 1980, The large-scale structure of the universe. Princeton Univ. Press, Princeton, NJ
- Pereira T., Pitrou C., 2015, Comptes rendus Physique, 16, 1027
- Perivolaropoulos L., Skara F., 2021, preprint (arXiv:2105.05208)
- Pezzotta A. et al., 2017, A&A, 604, A33
- Planck Collaboration, 2020, A&A, 641, A6
- Rasmussen C. E., 2003, Gaussian Processes in Machine Learning In Summer School on Machine Learning, 63. Springer, Berlin, Heidelberg.
- Renzi F., Silvestri A., 2020, preprint (arXiv:2011.10559)
- Řípa J., Shafieloo A., 2019, MNRAS, 486, 3027
- Ross N. P. et al., 2007, MNRAS, 381, 573
- Sagredo B., Nesseris S., Sapone D., 2018, Phys. Rev. D, 98, 083543
- Schneider P., 2006, Extragalactic Astronomy and Cosmology: An Introduction. Springer, Berlin
- Scrimgeour M. I., 2012, MNRAS, 425, 116
- Seikel M., Clarkson C., Smith M., 2012, J. Cosmol. Astropart. Phys., 06, 036
- Shafieloo A., Kim A. G., Linder E. V., 2012, phys. Rev. D, 85, 123530
- Silk J., 1968, ApJ, 151, 459
- Song Y. S., Percival W. J., 2009, J. Cosmol. Astropart. Phys., 10, 004
- Strauss M. A., Willick J. A., 1995, Phys. Rep., 261, 271
- Sun W., Jiao K., Zhang T. J., 2021, preprint (arXiv:2105.12618)
- Sunyaev R. A., Zeldovich Y. B., 1970, Ap&SS, 7, 3
- Tarnopolski M., 2017, MNRAS, 472, 4819
- Tegmark M. et al., 2006, Phys. Rev. D, 74, 123507
- Velasquez-Toribio A. M., Fabris J. C., 2020, Eur. Phys. J. C., 80, 1
- Zhang M. J., Li H., 2018, Eur. Phys. J. C, 78, 1

### APPENDIX A: BARYON ACOUSTIC OSCILLATIONS INFLUENCE ON THE DPL APPROXIMATION

In Section 4.2, we have seen a discrepancy of 0.3 per cent between the fiducial model and the DPL model for  $R_{\rm H} \simeq 80 \text{ Mpc} h^{-1}$ . This small deviation in the fit can be attributed to the Baryon Acoustic Oscillations (BAO) signature, present around the scale 100 Mpc  $h^{-1}$ . To test this hypothesis, we perform our fit for the DPL approximation considering two estimates of the correlation function: one from the CLASS code and another for the case of absence of the BAO feature. For the last case, we use the fitting model given by Eisenstein & Hu (1998) and implemented in the code nbodykit<sup>4</sup> to obtain  $\xi(r)$ .

Fig. A1 shows the relative difference  $\zeta_{\text{DPL}}/\zeta_{\text{ACDM}} - 1$  obtained calculating the correlation function with and without the BAO feature. It is evident the improvement in the fitting of  $\zeta$  obtained





**Figure A1.** The relative difference  $\zeta_{DPL}/\zeta_{\Lambda CDM} - 1$  obtained calculating the correlation function with and without the BAO feature.

from the correlation function without the BAO feature for  $R_{\rm H} < 40$  Mpc  $h^{-1}$ , with more significant effect for  $R_{\rm H} \gtrsim 55$  Mpc  $h^{-1}$  when compared to the correlation function with BAO.

### APPENDIX B: STUDYING THE PARAMETER DEPENDENCE OF THE GROWTH RATE

In Section 3.1, we found a relation between f(z) and  $R_H(z)$  that is, in principle, explicitly independent of the  $\epsilon$  parameter. However, when measuring  $R_H(z)$ , one needs to fix  $\epsilon$ . Then, it is important to check if this criterion affects the f(z) estimate. Here, we investigate the impact of fixing  $\epsilon$ , as well as whether our choice of cosmological parameters might affect the f(z) obtained.

We compare the result obtained from our fiducial model,  $f^{\text{fiducial}}(z)$ , using the input parameters

$$\{\epsilon, \ln(10^{10}A_s), \Omega_K, \Omega_c h^2\} = \{0.01, 3.045, 0.0, 0.1202\},$$
 (B1)

with the f(z) resulting from the same fitting procedure but now varying these four parameters one at a time. The comparison is performed through the relative difference  $f(z)/f^{\text{fiducial}}(z) - 1$ .

Fig. B1 displays the relative difference between our input model and the 3 cases studied where  $\epsilon = \{0.02, 0.05, 0.001\}$ . These values correspond to different definitions of the homogeneity scale,  $R_H$ , where this scale is obtained when the data in analysis reaches 2 per cent, 0.5 per cent, and 0.1 per cent below the limit value 3, respectively (Scrimgeour et al. 2012). We show that for the redshift interval of interest, 0 < z < 2, the error is below 0.5 per cent, which makes our approach robust with respect to  $\epsilon$ . For the case  $\epsilon = 0.001$ , a divergent behaviour is observed around  $z \sim 1.4$ , where the function



Figure B1. The relative difference for f(z) between our fiducial model, i.e.  $\epsilon = 0.01$ , and the cases investigated with  $\epsilon = \{0.02, 0.005, 0.001\}$ .



**Figure B2.** The relative difference,  $f(z)/f^{\text{fiducial}}(z) - 1$ , considering the variations, one at a time, of three cosmological parameters:  $\ln(10^{10}A_s)$ ,  $\Omega_K$ , and  $\Omega_c h^2$ . The left-hand panel shows the results for  $A_s$ , which is basically noise. The middle panel shows the outcomes for  $\Omega_K$  where the relative difference is less than 6 per cent for the whole interval of interest. The right-hand panel shows the dependence on  $\Omega_c h^2$  which is  $\lesssim 4$  per cent, with the largest values for low z.

explodes up and come back from below. However, we notice that such small  $\epsilon$  is unpractical when investigating  $R_H$  due to the statistical errors (and other systematics) inherent to the data analyses.

Fig. B2 shows  $f(z)/f^{fiducial}(z) - 1$  for analyses obtained with the variation of three cosmological parameters one at a time:  $\ln (10^{10}A_s)$ ,  $\Omega_K$ , and  $\Omega_c h^2$ . For  $\ln (10^{10}A_s)$  we use {2.9, 3.1, 3.2}, which is a large enough  $\Omega_c h^2$ . For  $\ln (10^{10}A_s)$  we use {2.9, 3.1, 3.2}, which is a large enough of Fig. B2, show nothing but statistical noise, indicating that our model is independent of  $\ln (10^{10}A_s)$  values. For  $\Omega_K$ , we consider { -0.1, -0.01, 0.1} (see Fig. B2, middle panel). Also well beyond  $2\sigma$  uncertainty for the Planck best fit  $\Omega_K = -0.044 \pm 0.0165$ . For all these cases, we observe a maximum of 6 per cent deviation, for the whole redshift interval. For all purposes, our approach has a small dependence on  $\Omega_K$  considering a large interval of possible values.

For the analyses of the last parameter,  $\Omega_c h^2$ , we consider {0.11, 0.125, 0.13}. In these cases, we also find a slight dependence in our results,  $\leq 4$  per cent, and decreasing for high z (see the right-hand panel of Fig. B2). As in the previous analyses, this result was somehow expected, because we are not modifying the meaning of f(z), we just found an alternative way to find it. We already knew that the growth rate has a strong dependence in the matter density parameter, as seen in the parametrization  $f(z) = \Omega_m(z)^{\gamma}$ , where  $\gamma$  depends only on the constant of the equation of state,  $\omega = -1$ , for the  $\Lambda$ CDM case, or modifications according to the gravity model used.

### APPENDIX C: CONSISTENCY TEST FOR DIFFERENT KERNELS

Our main result, i.e. the reconstruction of the homogeneity scale function, which, in turn, we use to derive the growth rate of structures, is based on the SE kernel. The SE kernel is a smooth covariance function that can reproduce global characteristics, although sometimes it cannot reproduce local characteristics.

As our data sample is nicely distributed, this kernel works smoothly. In order to test for possible systematic effects on the kernel choice, we examine our main results using this time the Matérn class kernels. The Matérn kernel can be written as

$$K_{M_{\nu}}(\tau) = \sigma_f^2 \frac{2^{1-\nu}}{\Gamma(\nu)} \left(\frac{\sqrt{2\nu} \tau}{l}\right)^{\nu} K_{\nu} \left(\frac{\sqrt{2\nu} \tau}{l}\right), \tag{C1}$$

where  $K_{\nu}$  is the modified Bessel function of second kind,  $\Gamma(\nu)$  is the standard Gamma function and  $\nu$  is a strictly positive parameter.



**Figure C1.** The reconstruction of the growth rate of structures  $f^{R_{\rm H}}(z)$  using two different kernels, namely, kernel 1 (the SE kernel) and kernel 2 (the Matérn kernel).

Here, the hyper-parameters  $\sigma_f$  and l are also optimized during the fitting.

Fig. C1 shows the best-fitting prediction and the GP reconstruction of the  $f^{R_{\rm H}}(z)$  function using the SE and the Matérn kernels. We do not find significant differences between both reconstructions; therefore, we conclude that they are statistically equivalent.

# APPENDIX D: MEASUREMENTS OF THE GROWTH RATE FUNCTION

The literature reports diverse compilations of measurements of the growth rate of cosmic structures, f(z) (see e.g. Basilakos 2012; Nunes et al. 2016; Sagredo et al. 2018), which we update here. Our compilation, shown in Table D1, follows three criteria in order to avoid or minimize possible data correlations, we consider: (i) f(z) measurements obtained with cosmic tracers from different astronomical surveys or from disjoint redshift bins; (ii) direct measurements of f, and not measurements of  $f\sigma_8$  that use a fiducial cosmological model to eliminate the  $\sigma_8$ -dependence; (iii) the latest measurement of f when the same astronomical survey performed two or more measurements corresponding to several data releases.

### 10 F. Avila et al.

Survey	Z	f	Reference	Cosmological tracer
ALFALFA	0.013	$0.56\pm0.07$	Avila et al. (2021)	H I extragalactic sources
2dFGRS	0.15	$0.49\pm0.14$	Hawkins et al. (2003), Guzzo et al. (2008)	Galaxies
GAMA	0.18	$0.49\pm0.12$	Blake et al. (2013)	Multiple-tracer: blue & red gals.
WiggleZ	0.22	$0.60\pm0.10$	Blake et al. (2011)	galaxies
SDSS	0.35	$0.70\pm0.18$	Tegmark et al. (2006)	Luminous red galaxies (LRG)
GAMA	0.38	$0.66\pm0.09$	Blake et al. (2013)	Multiple-tracer: blue & red gals.
WiggleZ	0.41	$0.70\pm0.07$	Blake et al. (2011)	Galaxies
2SLAQ	0.55	$0.75\pm0.18$	Ross et al. (2007)	LRG & QSO
WiggleZ	0.60	$0.73\pm0.07$	Blake et al. (2011)	Galaxies
VIMOS-VLT Deep Survey	0.77	$0.91\pm0.36$	Guzzo et al. (2008)	Faint galaxies
2QZ & 2SLAQ	1.40	$0.90\pm0.24$	Da Ângela et al. (2008)	QSO

**Table D1.** Data compilation of f(z) measurements that shares important features, as explained in the Appendix D.

This paper has been typeset from a  $T_{\ensuremath{E}}X/I\!\!\!\! \mbox{AT}_{\ensuremath{E}}X$  file prepared by the author.